Discrete Mathematics and Probability Week 8

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Topics

- \blacktriangleright Bernoulli distribution: single trial
- \blacktriangleright Binomial distribution: many independent trials
- **Poisson distribution: counting independent trials**
- Geometric distribution: first success in independent trials

[Bernoulli and Binomial distributions](#page-2-0)

Bernoulli distribution

Definition

Suppose that *n* independent trials are performed, each succeeding with probability *p*. Let *X* count the number of successes within the *n* trials. Then *X* has the *Binomial distribution with parameters n and* p or, in short, $X \sim Bin(n, p)$. I

Special case $n = 1$ is called *Bernoulli distribution with parameter p.*

Bernoulli: mass function

Proposition *Let* $X \sim Bin(n, p)$. Then $X = 0, 1, ..., n$, and its mass function is $p(i) = P(X = i) = {n \choose i}$ *i* ◆ $p^{i}(1-p)^{n-i},$ $i = 0, 1, ..., n.$

In particular, the Bernoulli(*p*) *variable can take on values* 0 *or* 1*, with respective probabilities*

$$
p(0) = 1 - p, \qquad p(1) = p.
$$
\n
$$
p(\text{not a six}) = 1 - \frac{1}{6} = \frac{5}{6}
$$
\n
$$
P(a \text{ six on } 1 \text{ only}) = \frac{1}{6}
$$

Mass function

$$
n = 1
$$
 trial with 1 *event*
\n
$$
P(x=1) = p = 1 - p(x=0) = 1 - p
$$
 (4)
\n
$$
f(k:p) = \begin{cases} p & \text{if } k=1 \text{ (success)} \\ 1-p = q & \text{if } k=0 \text{ (failure)} \end{cases}
$$
\n
$$
f(k:p) = p^{k} \cdot (1-p)^{1-k} \quad \text{for } k \in \{0,1\}
$$
\n
$$
Bernoulli \quad \text{diviribubin} \quad \text{if } a \quad \text{special } \quad \text{case } \quad \text{aff}
$$
\n
$$
the \quad Binomial \quad \text{diviribubin} \quad \text{with } n=1,
$$

 \bullet

Bernoulli: expectation, variance

Proposition $\textit{Let } X \sim B$ **in**(\bullet , p). Then: $E(X) = pp$, and $Var(X) = pp(1-p) = pq$ Nol 1 - p = 9 $X \sim$ Ber (p) $E(x) = P$. Var $(x) = P(1-P)$ $\mu = E(x) = \oint_{x} p(x) = \int_{0}^{x} P(x=1) + \bigcirc_{0} P(x=0)$ $= 1 \cdot p + 0 \cdot q = p$ $Var(X) = E(X^2) - \mu^2 = \rho - \rho^2 = PL^{1-p} = PV$
 $E(v_i) = C \rho^2 - \mu^2 = (m - 1) + \rho^2 \rho(v - \rho) = p$ $E(X^2) = \mathcal{E}(X)P(x) = \frac{1}{\pi}P(X=1) + O(P(X=0)) = P$

Proof

Example

Proposition

Let $X \sim Bin(n, p)$. Then $X = 0, 1, ..., n$, and its mass function is

$$
\mathfrak{p}(i) = \mathbf{P}(X = i) = \binom{n}{i} p^{i} (1-p)^{n-i}, \qquad i = 0, 1, ..., n.
$$

 $16/22$ Binomial variate can be considered as the sum of ⁿ independent Bernoulli variates with the probability of 'success'. $X \sim Bin(n, p) \equiv X \sim Ber(p) + Ber(p)$ Ber (p) n times $M_{\text{BIN}}(n, p) = E(X)$ = $\mu_{\text{ger}} + \mu_{\text{ger}} + \cdots$ μ_{ger} ntime σ^2 σ^2 ρ^2 ρ^2 Var_{θ} (n, e) = $VAR(x) + \cdots$ Var(x) (n tunes P \rightarrow P

Examples

Probability of obtaining 2 sixes on tossing a die 3 times

\n
$$
S = \begin{cases}\n\text{NNN, NMS, } \text{NSO, SNN, NSS,} \\
\text{SNS, } \text{SSN, SSM, NSS,} \\
\text{SNS, } \text{SSN, SSM}\n\end{cases}
$$
\n
$$
S = \begin{cases}\n\frac{1}{2}N\frac{1}{2}V\frac{1}{2} & \text{S.} \\
\frac{1}{2}N\frac{1}{2}V\frac{1}{2} & \text{S.} \\
\frac{1}{2}N\frac{1}{2}V\frac{1}{2}V\frac{1}{2} & \text{S.} \\
\frac{1}{2}N\frac{1}{2}V\frac{1}{2}V\frac{1}{2}V\frac{1}{2} & \text{S.} \\
\frac{1}{2}N\frac{1}{2}V\frac{1}{2}V\frac{1}{2}V\frac{1}{2} & \text{S.} \\
\frac{1}{2}N\frac{1}{2}V\frac{1}{2}V\frac{1}{2}V\frac{1}{2} & \text{S.} \\
\frac{1}{2}N\frac{1}{2}V\frac{1}{2}V\frac{1}{2}V\frac{1}{2} & \text{S.} \\
\frac{1}{2}N\frac{1}{2}V\frac{1}{2}V\frac{1}{2} & \text{S.} \\
\frac{1}{2}N\frac{1}{2}V\frac{1}{2}V\frac{1}{2}V\frac{1}{2} & \text{S.} \\
\frac{1}{2}N\frac{1}{2}V\frac{1}{2}V\frac{1}{2}V\frac{1}{2} & \text{S.} \\
\frac{1}{2}N
$$

[Poisson distribution](#page-11-0)

Poisson: mass function

The Poisson distribution is of central importance in Probability. Will later see relation to Binomial. of everts

Definition

Fix a positive real number λ . The random variable X is *Poisson distributed with parameter* λ , in short $\overline{X} \sim \text{Poi}(\lambda)$, if it is non-negative integer valued, and its mass function is

$$
p(i) = P(X = i) = e^{-\lambda} \cdot \frac{\lambda^i}{i!}, \qquad i = 0, 1, 2, ...
$$

For $X \sim Poi(\lambda)$, $E(X) = Var(X) = \lambda$.

is mean number

Poisson approximation of Binomial

Proposition

Fix $\lambda > 0$, and suppose that $Y_n \sim Bin(n, p)$ with $p = p(n)$ in such *a* way that $n \cdot p \rightarrow \lambda$. Then the distribution of Y_n converges to $Poisson(\lambda)$:

$$
\forall i \geq 0 \qquad \mathbf{P}(Y_n = i) \underset{n \to \infty}{\longrightarrow} e^{-\lambda} \frac{\lambda^i}{i!}.
$$

Example

Poisson: many independent small prob events.
\nsummin' of the "a few" in expectation.
\nBi
$$
(X=i)
$$
 $2x s$: no of types on a page of a book.
\nno of citizens & 100 yrs old in
\na city
\n $e^{-\lambda} \frac{\lambda}{i!}$ on o of calls per how in a customer
\n $0! \equiv 1$ on o of customers in P.0. today.
\nbook has a average of $\frac{1}{2}$ types per page.
\nchange that the next page has x 3 types
\n $X \sim Poi (A)$. $\lambda = Y_2$
\n $P(x \gg 3) = 1 - P(X \le 2)^2$
\n $= 1 \cdot 4^{\frac{5}{2}} \cdot 1 - \frac{P(x=0)}{2! \cdot 2!} - \frac{P(x=1)}{2! \cdot 2! \cdot 2!} - \frac{P(x=2)}{2! \cdot 1! \cdot 2!} = \frac{P(x=3)}{2! \cdot 1! \cdot 2!}$

Example

A Geiger counter is used in recording radioactive events. Each radioactive event arriving at the counter shows as a number. The number of radioactive events recorded in a room is on average 2 every 5 seconds - background count.

I

a) Find the probability of exactly 3 events in 5 seconds
\nb) The probability of more than 3 events in 5 seconds.
\nA = 2
\na)
$$
P_{oi} (x = 3) = 2^3 e^{-2}
$$

\nb) $1 - P_{gi} (0) - P_{gi} (1) - P_{oi} (2) - P_{oi} (3)$
\n $P_{oi} (x = 1) = 2^3 e^{-2}$
\n $3! - P_{oi} (2) - P_{oi} (3)$
\n $P_{oi} (4) - P_{oi} (5) - P_{oi} (6) - P_{oi} (7) - P_{oi} (8) - P_{oi} (9) - P_{oi} (1) - P$

[Geometric distribution](#page-16-0)

Geometric: mass function

Again independent trials, but now ask: when is the first success?

Definition

Suppose that independent trials, each succeeding with probability *p*, are repeated until the first success. The total number *X* of trials made has the *Geometric*(p) distribution (in short, $X \sim \text{Geo}(p)$).

Proposition

X can take on positive integers, with probabilities $p(i) = (1-p)^{i-1} \cdot p, i = 1, 2, \ldots$

$$
\frac{5}{6} > \frac{5}{6} \pi \frac{5}{6} \text{ s.t. } \frac{10}{3} \text{ s.t. } \frac{5}{6} \text{ s.t. } \frac{5}{6} \text{ s.t. } \frac{1}{6} \text{ s.t. } \frac{5}{6} \text{ s.t. } \frac{1}{6} \text{ s.t. }
$$

Geometric: mass function

Corollary

The Geometric random variable is (discrete) memoryless:

$$
\mathbf{P}\{X\geq n+k\,|\,X>n\}=\mathbf{P}\{X\geq k\}
$$

for every $k \geq 1$, $n \geq 0$.

Geometric: expectation, variance

Proposition *For a Geometric*(*p*) *random variable X:* $E(X) = \frac{1}{x}$ *p* $\mathsf{Var}(X) = \frac{1-p}{2}$ *p*2

Example

Summary

- \blacktriangleright Bernoulli distribution: single trial
- \blacktriangleright Binomial distribution: many independent trials
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