Foundations of Data Science: Regression and inference – From the maximum likelihood principle to linear regression



THE UNIVERSITY of EDINBURGH

EOUNDATIONS OF DATA SCIENCE We want to investigate the relationship between the number of bikes hired in a day and the temperature on that day

Is there a problem with using ordinary least squares linear regression to do this?

Data sources:

- Edinburgh Just Eat Bikes data 2020
- Edinburgh temperature observations, Met Office via MIDAS



Overview

Today

- 1. The maximum likelihood principle
- Application of maximum likelihood principle to a simple example
- 3. Application of maximum likelihood principle to linear regression

Wednesday

- Max likelihood with non-normal distributions
- Generalised linear regresion

Foundations of Data Science: Regression and inference – The maximum likelihood principle

Intuition for maximum likelihood principle

Statistic model $P(Y=y|2, i\cdot 1) = \frac{1}{2\pi}e^{-\frac{1}{2}(\frac{y-2}{1-1})}$ $\bigwedge_{\mathcal{M}} \int_{\sigma^2} \int_{\sigma^2} \frac{1}{2\pi}i\cdot 1$ e, q Normal dist. $-\frac{1}{2} + \frac{1}{2} + \frac{1$ "y is drawn Alternative Notation: $y \sim N(2, |\cdot|^2)$ $\uparrow \quad \mu \quad \backslash_{\sigma^2}$ From normal dist with me un 2 and variance 1.12" nomal dist

Intuition for maximum likelihood principle Statistical model $P(Y=y|M,\sigma^{2}) = \frac{1}{2\pi}e^{-\frac{1}{2}\left(\frac{y-\mu}{\sigma}\right)}$ $\bigwedge_{M} \int_{\sigma^{2}} \int_{\sigma^{2}} \frac{1}{2\pi}e^{-\frac{1}{2}\left(\frac{y-\mu}{\sigma}\right)}$ "data" Infer Mandoz Mandoz Find porameters (here μ , σ^2) that maximise likelihood of he model having produced the observed data.

Exercise

which model is Data most likely to, have generated, this data? 2 ٥ $(\mathbf{j}) \mathbf{y}_{i} \sim \mathcal{N}(\mathbf{0},\mathbf{1})$ yn Jar yn (1,0·1²) $y_i \sim \mathcal{N}(1, 5^2)$ 3 $Y_i \sim N(1,1)$ 4)

Definition of the maximum likelihood principle

For a set of observed data and a given statistical model the principle of maximum likelihood states that the parameters of the model are adjusted so as to maximise the likelihood that the model generated the observed data.

Foundations of Data Science: Regression and inference – Application of the maximum likelihood principle to a simple example

Application to 1-variable example

Assume samples are drawn independently
 Assume each sample is drawn from an normal distribution

$$P(\gamma = y_i|\mu, \sigma^2) = \frac{1}{12\pi}e^{-\frac{1}{2}(\frac{y_i-\mu}{\sigma^2})^2}$$

Assumption
$$(\mathcal{Y} = \mathcal{Y})$$

 $P(\mathcal{Y} = (\mathcal{Y}_1, \dots, \mathcal{Y}_n) | \mu, \sigma^2) = P(\mathcal{Y} = \mathcal{Y}_1 | \mu, \sigma^2) \times P(\mathcal{Y} = \mathcal{Y}_2 | \mu, \sigma^2) \times P(\mathcal{Y} = \mathcal{Y}_2$

$$P(Y=y, 1\mu, \sigma^2)$$

More compact notation...

$$P(Y = (y_1, \dots, y_n) | \mu_1, \sigma^2) = P(Y = y_1 | \mu_1, \sigma^2) \times P(Y = y_2 | \mu_1, \sigma^2) \times P(Y = y_1 | \mu_1, \sigma^2) \times P(Y = y_1 | \mu_1, \sigma^2)$$

$$= \prod_{i=1}^{n} P(Y = y_i | \mu_1, \sigma^2)$$

$$= \prod_{i=1}^{n} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}} \left(\frac{y_i - \mu_1}{\sigma^2} \right)^2$$

Likelihood as a function of parameters

Data: Y.s. yis drawn from $\mathcal{N}(0,1)$ -1.00 × -0.75 -0.50-0.25 0.00 ~ 0.25 0.50 0.75 0 [(ode]



Log-likelihood as a function of parameters



Log-likelihood equations: products to sums

$$ln \alpha b = ln \alpha + ln b$$
So
$$ln(p_{1} \times p_{2} \times \dots \times p_{n}) = lnp_{1} + lnp_{2} + \dots + lnp_{n}$$
So
$$ln \prod_{i=1}^{n} p_{i} = \sum_{i=1}^{n} ln p_{i}$$

$$ln \prod_{i=1}^{n} P(\gamma = y_{i}) \mu_{i} \sigma^{2}) = \sum_{i=1}^{n} ln P(\gamma = y_{i}) \mu_{i} \sigma^{2}$$

The beauty of logs and sums

- Sum of logs is easy to represent within limits of floating point arithmetic
- Log likelihood function is smoother than likelihood function
- Sums are easy to differentiate; products are not

The log of the normal distribution

$$\ln \int \frac{1}{\sqrt{12\pi}} e^{-\frac{1}{2}} \left(\frac{y-r}{\sigma}\right)^2 \int$$

$$= - \ln \left(\sqrt{2\pi} \sigma \right) - \frac{1}{2} \left(\frac{y - \mu}{\sigma} \right)^{2}$$
$$= - \frac{1}{2} \ln 2\pi \sigma^{2} - \frac{1}{2} \left(\frac{y - \mu}{\sigma} \right)^{2}$$

Final expression for log likelihood

$$ln \quad p\left(\underbrace{Y}=y_{1}, \ldots, y_{n} | \mu_{1}\sigma^{2}\right)$$
$$= \underbrace{\widehat{\mathcal{C}}}_{i=1}\left(-\underbrace{l} \ln 2\pi\sigma^{2} - \frac{1}{2}\left(\underbrace{Y_{i}-\mu}_{\sigma}\right)^{2}\right)$$

Maximise w.r. $t \mu$ and σ^{2} => Maximum like lihood estimates (MLE) $\hat{\mu} = \frac{1}{n} \sum_{i=1}^{n} y_{i}$ $\hat{\sigma}^{2} = \frac{1}{n} \sum_{i=1}^{n} (y_{i} - \hat{\mu})^{2}$

Exercise: prove these statements by differentiating
w.r.t.
$$\mu$$
 and σ^2

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Relationship to ordinary least squares

$$\ln p(Y = y_{1}, \dots, y_{n}; x_{1}, \dots, x_{n} | B_{0}, B_{1}, \sigma^{2})$$

$$= \sum_{i=1}^{n} \left(-\frac{1}{2} \ln 2\pi \sigma^{2} - \frac{1}{2} \left(\frac{y_{i} - B_{0}}{\sigma^{2}} - B_{0} x_{i} \right)^{2} \right)$$

$$= SSE$$

Estimates of coefficients

Analytical solutions for
$$\hat{B}_0$$
, \hat{B}_1 , $\hat{\sigma}^2$ that maximise likelihing
 \hat{B}_0 and \hat{B}_1 , as per ordinary least squares
Variance of residuals:
 $\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^{2^*} (y_i - \frac{\hat{B}_0 - \hat{B}_1 2i}{n})^2$
 $= \frac{1}{n} \sum_{i=1}^{2^*} (y_i - \frac{\hat{y}_i}{2})^2 = \frac{SSE}{n} = \frac{-Biased}{n}$
Sumpling theorem $\hat{\sigma}^2 = \frac{SSE}{n-2} \in \text{Unbiassed}$

Log likelihood of coefficients



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Data from Wauters and Dhondt 1989

Bootstrap inference of coefficients



(Log) Likelihood function

Bootstrap samples



Overview

- 1. Maximum likelihood principle
 - What model was most likely to have generated the data
- 2. Maximum likelihood principle applied to simple example
 - Log likelihood turns out to be useful
 - Gives rise to familiar estimates for mean and variance
- 3. Maximum likelihood principle applied to linear regression
 - Turns out to give ordinary least squares
 - Link with coefficient uncertainty and the bootstrap estimates of parameter uncertainty

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