# Inf2 - Foundations of Data Science 2023 <br> Workshop solution: Semester 2 Week 6 Workshop 

8th February 2024

## 1. Logistic Regression

(a) $\hat{\beta}_{0}=1.0$. More likely to be rejected, since the $\log$ odds are above 0 .
(b) $\exp \left(\hat{\beta}_{0}\right)=2.718$
(c) $p$ (Reject $)=\frac{1}{1+\exp \left(-\hat{\beta}_{0}\right)}=0.731$ probability of rejection
(d) Odds ratio

$$
\begin{aligned}
\log \text { odds } & =\hat{\beta}_{0}+\hat{\beta}_{1} x^{(1)}+\hat{\beta}_{2} x^{(2)}+\hat{\beta}_{3} x^{(3)} \\
& =1.0+0.5 \times 5-0.5 \times 0-0.1 \times 0 \\
& =3.5
\end{aligned}
$$

$$
\begin{aligned}
\text { Odds }(\text { Reject }) & =\exp \left(\hat{\beta}_{0}+\hat{\beta}_{1} x^{(1)}+\hat{\beta}_{2} x^{(2)}+\hat{\beta}_{3} x^{(3)}\right) \\
& =\exp (3.5)=33.1
\end{aligned}
$$

$$
\begin{aligned}
p(\text { Reject }) & =\frac{1}{1+\exp \left(-\left(\hat{\beta}_{0}+\hat{\beta}_{1} x^{(1)}+\hat{\beta}_{2} x^{(2)}+\hat{\beta}_{3} x^{(3)}\right)\right)} \\
& =1 /(1+\exp (-3.5))=0.971
\end{aligned}
$$

(e)

$$
\begin{aligned}
p(\text { Reject }) & =\frac{1}{1+\exp \left(-\left(\hat{\beta}_{0}+\hat{\beta}_{1} x^{(1)}+\hat{\beta}_{2} x^{(2)}+\hat{\beta}_{3} x^{(3)}\right)\right)} \\
& =\frac{1}{1+\exp (-(1+0.5 \times 0-0.5 \times 3-0.1 \times 2))} \\
& =\frac{1}{1+\exp (0.7)}=0.332
\end{aligned}
$$

(f)

$$
\begin{aligned}
\text { Log odds } & =\hat{\beta}_{0}+\hat{\beta}_{1} x^{(1)}+\hat{\beta}_{2} x^{(2)}+\hat{\beta}_{3} x^{(3)} \\
& =1+0.5 \times 1-0.5 \times 3-0.1 \times 1 \\
& =-0.1<0
\end{aligned}
$$

Since the log odds are less than 0 (which corresponds to probability of rejection equal to 0.5 ), we do not reject the paper.
(g) The probability of rejection equal to 0.25 is equal to $\log$ odds of $\ln 0.25 /(1-0.25)=$ -1.097 , so this is the threshold. The log odds computed in the previous part now exceed this threshold, so the paper will be rejected.
(h) You could explain that when you used to review papers yourself a paper that does not contain any of the phrases "world-beating", "confidence interval" or "bootstrap" would have had a probability of 0.731 of being rejected, or, in other words, it was 2.718 times more likely to be rejected than accepted. You could then say that every extra occurrence of the word "world-beating" increased the odds of rejection by 1.65 times (i.e. $e^{\hat{\beta}_{1}}$ ), but that the word "confidence interval" reduced the odds by a factor of $1.65\left(e^{\hat{\beta}_{2}}\right)$, and the word "bootstrap" reduced the odds by a factor of 1.11 $\left(e^{\hat{\beta}_{3}}\right)$.
Alternatively, you could say that you've now implemented a scoring system that is implemented by weighting the number of occurrences of each word, and give the weights of each word and the threshold, which is $\ln 1 / 3-\hat{\beta}_{0}=-2.1$, assuming a probability threshold of 0.25 , i.e. an odds threshold of $1 / 3$.

## 2. $A / B$ testing

(a) Let $p_{1}$ denote the proportion who responded to the sun lounger picture, and let $p_{2}$ who responded to the beach filled with people.
The sample estimates of the true proportions are $\hat{p_{1}}=224 / 500=0.448$ and $\hat{p_{2}}=150 / 500=0.3$.
The estimator of the difference between the sample proportions is $\hat{d}=\hat{p_{1}}-\hat{p_{2}}=$ 0.148 .

The estimated standard error of the difference is

$$
\begin{aligned}
\hat{\sigma}_{\hat{d}} & =\sqrt{\frac{p_{1}\left(1-p_{1}\right)+p_{2}\left(1-p_{2}\right)}{n}} \\
& =\sqrt{\frac{0.448 \times(1-0.448)}{500}+\frac{0.3 \times(1-0.3)}{500}} \\
& =0.0302
\end{aligned}
$$

Assuming a $95 \% \mathrm{Cl}$ and using a normal approximation to the binomial we compute the two-sided confidence interval with $\alpha=0.05$ as:

$$
\hat{d} \pm z_{\alpha / 2} \hat{\sigma}_{\hat{d}}=0.148 \pm 1.96 \times 0.0302=0.148 \pm 0.0593=(0.0887,0.2073)
$$

The $95 \%$ confidence interval does not contain 0 , hence we can conclude that the sample proportions are sufficiently different and that the campaign with the sun lounger picture is more successful.
(b) It might be that the time of day that you ran the initial trial had a different demographic online than for the week as a whole. Or perhaps people in the UK do not represent decisions made worldwide.

## 3. Hypothesis testing

(a) $\mathrm{H}_{0}$ : The 42 out of 262 trades in which Dream received an Ender Perl arose from each trade having a probability of $4.73 \%$ of returning an Ender Perl.
$\mathrm{H}_{\mathrm{a}}$ : The trades occurred via cheating which made it more likely that Dream received Ender Pearls.
(b) The distribution implied by the null hypothesis is a binomial distribution with $p=0.0473$ and $n=262$ trials. As $n$ is large we can approximate it by a normal distribution with $\mu=n p=12.3926$ and $\sigma^{2}=n p(1-p)=11.8064$, so the standard deviation is 3.4360 . We should do an upper tailed test, since the alternative hypothesis suggests that the process returns more Ender Pearls. The value of $z$ is

$$
z=\frac{42-\mu}{\sigma}=8.6170
$$

To find the area in the upper tail, we need to compute $1-\Phi(z)=1-\Phi(8.6170)$. A value of $z$ this large isn't to be found in statistical tables. However, the scipy function scipy.stats.norm.sf() is equal to $1-\Phi(z)=1-\Phi(8.6170)$, so we compute:
scipy.stats.norm.sf (8.6170)
in python. The result is $3.446 \times 10^{-18}$, which is the chance that if the null hypothesis were true, 42 or more Ender Pearls from 262 trades with Piglins would result.
(c) With the binomial distribution $b(x ; n, p)$, we are looking for the number of successful trades $X$ to be greater than or equal to 42, i.e. $P(X \geq 42 ; n, p)=\sum_{x=42}^{n} b(x ; n, p)$. This is equivalent to $1-\sum_{x=0}^{41} b(x ; n, p)$, which is one minus the cumulative distribution function for the binomial distribution, $B(X \leq 41 ; n, p)$. The scipy function scipy.stats.binom.sf() is exactly $1-B(X \leq 41 ; n, p)$, i.e. 1 minus the cumulative distribution function. We compute the value of $1-B(X \leq 41 ; n, p)$ in Python like this:
scipy.stats.binom.sf(41, n, p)
This returns $5.65 \times 10^{-12}$, the probability of 42 or more out of 262 trades being successful under the null hypothesis. This probability is about $1.6 \times 10^{6}$ times higher than the normal approximation, but still very low.
(d) With the binomial distribution, we follow the pattern above and calculate 1 -$B(211-1 ; 305,0.5)=8.8 \times 10^{-12}$.

[^0]
[^0]:    ${ }^{1}$ The "sf" stands for "survival function".

