# Informatics 2D: Reasoning and Agents

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#### Lecture 25: Approximate inference in Bayesian Networks

#### Where are we?

Last time ...

- Inference in Bayesian Networks
- Exact methods: enumeration, variable elimination algorithm
- Computationally intractable in the worst case

Today ...

• Approximate Inference in Bayesian Networks

## Approximate inference in BNs

- Exact inference computationally very hard
- Approximate methods important, here randomised sampling algorithms
- Monte Carlo algorithms
- We will talk about two types of MC algorithms:
  - Direct sampling methods
  - 2 Markov chain sampling

Rejection sampling

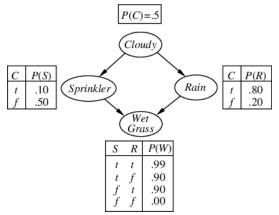
# Direct sampling methods

- Basic idea: generate samples from a known probability distribution
- Consider an unbiased coin as a random variable sampling from the distribution is like flipping the coin
- It is possible to sample any distribution on a single variable given a set of random numbers from [0,1]
- Simplest method: generate events from network without evidence
  - Sample each variable in 'topological order'
  - Probability distribution for sampled value is conditioned on values assigned to parents

Rejection sampling

## Example

• Consider the following BN and ordering [Cloudy, Sprinkler, Rain, WetGrass]:



#### Direct sampling methods

Likelihood weighting Inference by Markov chain simulation Summary

Example

- Direct sampling process:
  - Sample from P(Cloudy) = (0.5, 0.5), suppose this returns true
  - Sample from P(*Sprinkler*|*Cloudy* = *true*) = (0.1,0.9), suppose this returns *false*
  - Sample from  $P(Rain|Cloudy = true) = \langle 0.8, 0.2 \rangle$ , suppose this returns *true*
  - Sample from P(*WetGrass*|*Sprinkler* = false, *Rain* = true) = (0.9, 0.1), suppose this returns *true*
- Event returned=[true, false, true, true]

Rejection sampling

# Direct sampling methods

• Generates samples with probability  $S(x_1, \ldots, x_n)$ 

$$S(x_1,...,x_n) = P(x_1,...,x_n) = \prod_{i=1}^n P(x_i | parents(X_i))$$

i.e. in accordance with the distribution

- Answers are computed by counting the number  $N(x_1,...,x_n)$  of the times event  $x_1,...,x_n$  was generated and dividing by total number N of all samples
- In the limit, we should get

$$\lim_{n\to\infty}\frac{N(x_1,\ldots,x_n)}{N}=S(x_1,\ldots,x_n)=P(x_1,\ldots,x_n)$$

If the estimated probability P̂ becomes exact in the limit we call the estimate consistent and we write "≈" in this sense, e.g.
P(x<sub>1</sub>,...,x<sub>n</sub>) ≈ N(x<sub>1</sub>,...,x<sub>n</sub>)/N

Rejection sampling

## Rejection sampling

- Purpose: to produce samples for hard-to-sample distribution from easy-to-sample distribution
- To determine P(X|e) generate samples from the prior distribution specified by the BN first
- Then reject those that do not match the evidence
- The estimate  $\hat{P}(X = x | e)$  is obtained by counting how often X = x occurs in the remaining samples
- Rejection sampling is consistent because, by definition:

$$\hat{P}(X|\mathsf{e}) = rac{\mathsf{N}(X,\mathsf{e})}{\mathsf{N}(\mathsf{e})} \approx rac{\mathsf{P}(X,\mathsf{e})}{\mathsf{P}(\mathsf{e})} = \mathsf{P}(X|\mathsf{e})$$

Rejection sampling

#### Back to our example

- Assume we want to estimate P(Rain|Sprinkler = true), using 100 samples
  - 73 have Sprinkler = false (rejected), 27 have Sprinkler = true
  - Of these 27, 8 have Rain = true and 19 have Rain = false
- $P(\textit{Rain}|\textit{Sprinkler} = true) \approx \alpha \langle 8, 19 \rangle = \langle 0.296, 0.704 \rangle$
- True answer would be  $\langle 0.3, 0.7 \rangle$
- But the procedure rejects too many samples that are not consistent with e (exponential in number of variables)
- Not really usable (similar to naively estimating conditional probabilities from observation)

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# Likelihood weighting

• A direct sampling method that avoids inefficiency of rejection sampling,

by generating only samples consistent with evidence

- Fixes the values for evidence variables E and samples only the remaining variables X and Y
- Since not all events are equally probable, each event has to be weighted by its **likelihood** that it accords to the evidence
- Likelihood is measured by product of conditional probabilities for each evidence variable, given its parents

# Likelihood weighting

- Consider query P(*Rain*|*Sprinkler* = *true*, *WetGrass* = *true*) in our example; initially set weight *w* = 1, then event is generated:
  - Sample from  $P(Cloudy) = \langle 0.5, 0.5 \rangle$ , suppose this returns *true*
  - Sprinkler is evidence variable with value true, we set

 $w \leftarrow w \times P(Sprinkler = true | Cloudy = true) = 0.1$ 

- Sample from  $P(Rain|Cloudy = true) = \langle 0.8, 0.2 \rangle$ , suppose this returns *true*
- WetGrass is evidence variable with value true, we set

 $w \leftarrow w \times P(WetGrass = true | Sprinkler = true, Rain = true) = 0.099$ 

• Sample returned=[*true*, *true*, *true*] with weight 0.099 tallied under *Rain* = *true* 

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# Likelihood weighting – why it works

- $S(z,e) = \prod_{i=1}^{l} P(z_i | parents(Z_i))$
- S's sample values for each Z<sub>i</sub> is influenced by the evidence among Z<sub>i</sub>'s ancestors
- But *S* pays no attention when sampling *Z<sub>i</sub>*'s value to evidence from *Z<sub>i</sub>*'s non-ancestors; so it's not sampling from the true posterior probability distribution!
- But the likelihood weight *w* makes up for the difference between the actual and desired sampling distributions:

$$w(z,e) = \prod_{i=1}^{m} P(e_i | parents(E_i))$$

Likelihood weighting – why it works

• Since two products cover all the variables in the network, we can write

$$P(z, e) = \underbrace{\prod_{i=1}^{l} P(z_i | parents(Z_i))}_{S(z, e)} \underbrace{\prod_{i=1}^{m} P(e_i | parents(E_i))}_{w(z, e)}$$

- With this, it is easy to derive that likelihood weighting is consistent (tutorial exercise)
- Problem: most samples will have very small weights as the number of evidence variables increases

# The Markov chain Monte Carlo (MCMC) algorithm

- MCMC algorithm: create an event from a previous event, rather than generate all events from scratch
- Helpful to think of the BN as having a **current state** specifying a value for each variable
- Consecutive state is generated by sampling a value for one of the non-evidence variables X<sub>i</sub> conditioned on the current values of variables in the Markov blanket of X<sub>i</sub>
- Recall that Markov blanket consists of parents, children, and children's parents
- Algorithm randomly wanders around state space flipping one variable at a time and keeping evidence variables fixed

# The MCMC algorithm

- Consider query P(*Rain*|*Sprinkler* = *true*, *WetGrass* = *true*) once more
- Sprinkler and WetGrass (evidence variables) are fixed to their observed values, hidden variables Cloudy and Rain are initialised randomly (e.g. true and false)
- Initial state is [true, true, false, true]
- Execute repeatedly:
  - Sample *Cloudy* given values of Markov blanket, i.e. sample from P(*Cloudy*|*Sprinkler* = *true*, *Rain* = *false*)
  - Suppose result is *false*, new state is [*false*, *true*, *false*, *true*]
  - Sample *Rain* given values of Markov blanket, i.e. sample from P(*Rain*|*Sprinkler* = *true*, *Cloudy* = *false*, *WetGrass* = *true*)
  - Suppose we obtain *Rain* = *true*, new state [*false*, *true*, *true*, *true*]

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## The MCMC algorithm – why it works

- Each state is a sample, contributes to estimate of query variable *Rain* (count samples to compute estimate as before)
- Basic idea of proof that MCMC is consistent: The sampling process settles into a "dynamic equilibrium" in which the long-term fraction of time spent in each state is exactly proportional to its posterior probability
- MCMC is a very powerful method used for all kinds of things involving probabilities

# Summary

- Approximate inference in BN's
- Direct sampling methods
- Likelihood working and why it works
- MCMC algorithm and why it works
- Next time: Time and Uncertainty I