# Randomized Algorithms <br> Lecture 5 

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RA - Lecture 5 - slide 1

## Continuing review of Discrete Probability ... and the Coupon Collector Problem

> Recall: "Coupon collecting": we are buying boxes (of cereal), each of which has a uniformly random coupon inside. There are n different types of coupons, and the goal is to collect one of each type, and then stop buying.

Last time we showed the expected number $\mathrm{E}[X]$ of boxes we must buy is $n H(n)=n \ln (n)+\theta(n)$, or more precisely, $n \ln (n)+\frac{n}{2} \leq n H(n) \leq n \ln (n)+n$.

Today we examine what the probability is that a "run" of the purchasing process is far from that expectation.

Concentration inequalities will be vital:

- Markov Inequality;
- Chebyshev Inequality;
- Chernoff Bound / Hoeffding inequality.
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## Markov Inequality

Very simple and easy, but very important.
Theorem (3.1, Markov Inequality)
Let $X$ be any random variable that takes only non-negative values. Then for any $a>0$,

$$
\operatorname{Pr}[X \geq a] \leq \frac{E[X]}{a}
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Proof.
Define the indicator function $I=I(X)$ by

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I(x)= \begin{cases}0 & x<a \\ 1 & x \geq a\end{cases}
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Then $X \geq a \cdot I(X)$, and hence $I(X) \leq \frac{X}{a}$.
Taking expectation of both sides, and using $\mathrm{E}[I]=\operatorname{Pr}[X \geq a]$, we have

$$
\mathrm{E}[I]=\operatorname{Pr}[X \geq a] \leq \frac{\mathrm{E}[X]}{a}
$$

## Bounding Coupon Collector purchases - Markov

Let $X$ be the number of purchases we have to make in the coupon collector problem until we get all $n$ coupons. Recall, we know $\mathrm{E}[X]=n H_{n}$, and thus: $n \ln (n)+\frac{n}{2} \leq \mathrm{E}[X] \leq n \ln n+n$.

Suppose we want a lower bound $t$ on the number of boxes we have to buy, such that $\operatorname{Pr}[X \geq t] \leq \frac{1}{2}$.
By Markov's inequality, $\operatorname{Pr}[X \geq t] \leq \frac{\mathrm{E}[X]}{t} \leq \frac{n H_{n}}{t}$.
Thus, it suffices to let $t=2 n H_{n}$, to get $\operatorname{Pr}\left[X \geq 2 n H_{n}\right] \leq \frac{1}{2}$.
So, $\operatorname{Pr}[X \geq 2(n \ln (n)+n)] \leq \frac{1}{2}$.
However, this bound is way too weak: we can get far smaller probability of failure with $2 n H_{n}$ purchases.

The power of Markov ineq. is that it does not require any other knowledge of the random variable. However for specific problems, we can often do much better.

For example, we can bound the variance.

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## Variance and Covariance

Recall, the variance of a random variable is

$$
\operatorname{Var}[X]:=\mathrm{E}\left[(X-\mathrm{E}[X])^{2}\right]=\mathrm{E}\left[X^{2}\right]-\mathrm{E}[X]^{2} .
$$

Definition (3.3)
The covariance of two random variables $X$ and $Y$ is defined as

$$
\operatorname{Cov}[X, Y] \doteq \mathrm{E}[(X-\mathrm{E}[X])(Y-\mathrm{E}[Y])]=\mathrm{E}[X Y]-E[X] E[Y] .
$$

Theorem (3.2)
For any two random variables $X, Y$, we have

$$
\operatorname{Var}[X+Y]=\operatorname{Var}[X]+\operatorname{Var}[Y]+2 \operatorname{Cov}[X, Y] .
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Proof.

$$
\begin{aligned}
\operatorname{Var}[X+Y] & =\mathrm{E}\left[(X+Y)^{2}\right]-\mathrm{E}[X+Y]^{2} \\
& =\mathrm{E}\left[X^{2}\right]+\mathrm{E}\left[Y^{2}\right]+2 \mathrm{E}[X Y]-\mathrm{E}[X]^{2}-\mathrm{E}[Y]^{2}-2 \mathrm{E}[X] \mathrm{E}[Y] \\
& =\operatorname{Var}[X]+\operatorname{Var}[Y]+2(\mathrm{E}[X Y]-\mathrm{E}[X] \mathrm{E}[Y])
\end{aligned}
$$

## (pairwise) Independent Random Variables

Theorem (3.3)
If $X, Y$ are a pair of independent random variables, then

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Corollary (3.4)
If $X, Y$ are a pair of independent random variables, then

$$
\operatorname{Cov}[X, Y]=0
$$

and

$$
\operatorname{Var}[X+Y]=\operatorname{Var}[X]+\operatorname{Var}[Y] .
$$

## Chebyshev Inequality

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For every $a>0$,

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\operatorname{Pr}[|X-E[X]| \geq a] \leq \frac{\operatorname{Var}[X]}{a^{2}}
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Proof.
First note that for any $a>0$,

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|X-\mathrm{E}[X]| \geq a \Longleftrightarrow(X-\mathrm{E}[X])^{2} \geq a^{2}
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Apply Markov's Inequality to the random variable $(X-\mathrm{E}[X])^{2}$. We get:

$$
\operatorname{Pr}[|X-\mathrm{E}[X]| \geq a]=\operatorname{Pr}\left[(X-\mathrm{E}[X])^{2} \geq a^{2}\right] \leq \frac{\mathrm{E}\left[(X-\mathrm{E}[X])^{2}\right]}{a^{2}}=\frac{\operatorname{Var}[X]}{a^{2}}
$$

## Bounding Coupon Collector purchases - Markov

Recall that $X$ is the number of purchases of the coupon collector problem and $\mathrm{E}[X]=n H_{n} \leq n \ln n+n$.

Using Markov's inequality, we can get that $\operatorname{Pr}\left[X \geq n^{2} H_{n}\right] \leq \frac{1}{n}$.
We can do better with Chebyshev's inequality ...

$$
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$$

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- Need to evaluate $\operatorname{Var}[X]$, which is $\operatorname{Var}\left[X_{1}+\ldots+X_{n}\right]$.

Recall that $X_{i}$ is the number of boxes bought to get the $i$-th new card.

- Corollary 3.4: for independent $Y, Z, \operatorname{Var}[Y+Z]=\operatorname{Var}[Y]+\operatorname{Var}[Z]$.
- Are these $X_{i}$ 's independent? Yes.


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- Are these $X_{i}$ 's independent? Yes. $X_{i}$ is a geometrically distributed r.v. that only depends on the values $n$ and $i$ (and not on what cards we have collected or how long it took to collect them).
- Hence the random variables $X_{1}, \ldots, X_{n}$ are all mutually independent, and

$$
\operatorname{Var}[X]=\operatorname{Var}\left[X_{1}\right]+\operatorname{Var}\left[X_{2}\right]+\ldots+\operatorname{Var}\left[X_{n}\right] .
$$

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## Bounding Coupon Collector purchases - Chebyshev

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## Fact (3.8)

For any geometric random variable $X$ with parameter $p$,

$$
E[X]=p^{-1}, \quad \text { and } \quad \operatorname{Var}[X]=\frac{1-p}{p^{2}}
$$

(These facts are well known. See chapter 3 of book for a proof.)

$$
\operatorname{Pr}[|X-\mathrm{E}[X]| \geq a] \leq \frac{\operatorname{Var}[X]}{a^{2}}=\frac{\sum_{j=1}^{n} \operatorname{Var}\left[X_{j}\right]}{a^{2}}
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$$

Each individual $X_{j}$ is geometric with parameter $\frac{n-(j-1)}{n}$, so each $X_{j}$ has

$$
\operatorname{Var}\left[X_{j}\right]=\frac{j-1}{n}\left(\frac{n}{(n+1-j)}\right)^{2} \leq\left(\frac{n}{n+1-j}\right)^{2}
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\begin{gathered}
\operatorname{Var}\left[X_{j}\right]=\frac{j-1}{n}\left(\frac{n}{(n+1-j)}\right)^{2} \leq\left(\frac{n}{n+1-j}\right)^{2} \\
\operatorname{Var}[X] \leq n^{2} \sum_{j=1}^{n}\left(\frac{1}{n+1-j}\right)^{2}=n^{2} \sum_{j=1}^{n}\left(\frac{1}{j}\right)^{2}
\end{gathered}
$$

$$
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$$

## Bounding Coupon Collector purchases - Chebyshev

Theorem [Euler, 1741]
Hence,

$$
\sum_{i=1}^{\infty} \frac{1}{i^{2}}=\frac{\pi^{2}}{6}
$$

$$
\operatorname{Var}[X] \leq n^{2} \sum_{j=1}^{n} \frac{1}{j^{2}} \leq \frac{n^{2} \pi^{2}}{6}
$$

Suppose we are willing to make $2 \mathrm{E}[X]=2 n H_{n}$ purchases. The probability we fail to get all cards is

$$
\begin{aligned}
\operatorname{Pr}[X>2 \mathrm{E}[X]] & =\operatorname{Pr}[X-\mathrm{E}[X]>\mathrm{E}[X]] \\
& =\operatorname{Pr}[|X-\mathrm{E}[X]|>\mathrm{E}[X]] .
\end{aligned}
$$

$$
\text { (as } X \geq 0 \text { ) }
$$

## Bounding Coupon Collector purchases - Chebyshev

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\begin{align*}
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& =\operatorname{Pr}[|X-\mathrm{E}[X]|>\mathrm{E}[X]] .
\end{align*}
$$

Using Chebyshev Inequality with $a=\mathrm{E}[X]$ :

$$
\begin{aligned}
\operatorname{Pr}[|X-\mathrm{E}[X]| \geq \mathrm{E}[X]] & \leq \frac{\operatorname{Var}[X]}{\mathrm{E}[X]^{2}} \leq \frac{\pi^{2} n^{2}}{6 n^{2} H_{n}^{2}} \\
& =\frac{\pi^{2}}{6 H_{n}^{2}} \leq \frac{2}{\ln ^{2} n} .
\end{aligned}
$$

Note: This improves over $\frac{1}{2}$, which is what Markov gives us.

$$
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$$

## Bounding Coupon Collector purchases - Union bound

Theorem (1.2, Union bound)
Let $E_{1}, E_{2}, \ldots$ be a finite or countably infinite sequence of events. Then

$$
\operatorname{Pr}\left[\bigcup_{i \geq 1} E_{i}\right] \leq \sum_{i \geq 1} \operatorname{Pr}\left[E_{i}\right]
$$

Similar to Markov ineq., there is almost no requirement to the union bound!

## Bounding Coupon Collector purchases - Union bound

Let $E_{i}$ be the "bad" event where card $i$ is still missing at time $T$.

$$
\operatorname{Pr}\left[E_{i}\right] \leq\left(1-\frac{1}{n}\right)^{T}
$$

Thus, by a union bound,

$$
\operatorname{Pr}[X \geq T]=\operatorname{Pr}\left[\cup_{i \geq 1}^{n} E_{i}\right] \leq n\left(1-\frac{1}{n}\right)^{T}
$$

## Bounding Coupon Collector purchases - Union bound

Once again we use $(1-1 / n)^{n} \leq 1 / e$. If $T=(1+\varepsilon) n \ln n$,

$$
\begin{aligned}
n\left(1-\frac{1}{n}\right)^{T} & \leq n\left(\left(1-\frac{1}{n}\right)^{n}\right)^{(1+\varepsilon) \ln n} \\
& \leq n\left(e^{-1}\right)^{(1+\varepsilon) \ln n}=n^{-\varepsilon}
\end{aligned}
$$

Thus, for example if $\varepsilon=1$,

$$
\operatorname{Pr}[X \geq 2 n \ln n] \leq n^{-1} .
$$

As $\mathrm{E}[X] \geq n \ln n$,

$$
\operatorname{Pr}[X \geq 2 \mathrm{E}[X]] \leq \operatorname{Pr}[X \geq 2 n \ln n] \leq n^{-1} .
$$

## Coupon collector bounds

$$
\begin{align*}
& \operatorname{Pr}[X \geq 2 \mathrm{E}[X]] \leq \frac{1}{2}  \tag{Markov}\\
& \operatorname{Pr}[X \geq 2 \mathrm{E}[X]] \leq \frac{2}{\ln (n)^{2}}  \tag{Chebyshev}\\
& \operatorname{Pr}[X \geq 2 \mathrm{E}[X]] \leq \frac{1}{n}
\end{align*}
$$

(Union bound)

Using "Chernoff bounds" for "negatively correlated" r.v.'s, one can also show

$$
\operatorname{Pr}[X \leq(1-\varepsilon)(n-1) \ln n] \leq e^{-n^{\varepsilon}} .
$$

However, we will not establish this result in this course.

## Wrapping up today

Next week we will continue the theme of "bounding deviation from the mean" by introducing some very important concentration inequalities, which apply first and foremost to sums of independent random variables, called

Chernoff bounds / Hoeffding's inequality.
First, in the next lecture we give a simple randomized algorithm to approximate the Maximum Cut in a graph, and show how to derandomize it using conditional expectation.

